# Portfolio choice theory and asset pricing Tutorials : Risk measures

# Exercise 1

Assume that the loss distribution of a portfolio follows a Gaussian law  $L \sim \mathcal{N}(\mu, \sigma^2)$ . Then, prove that

$$\operatorname{Var}_{\alpha}(L) = \mu + \sigma \Phi^{-1}(\alpha),$$

where  $\Phi$  is the cdf of  $\mathcal{N}(0,1)$  and  $\Phi^{-1}(\alpha)$  is the  $\alpha$ -quantile of  $\Phi$ .

## Exercise 2

We consider a portfolio consisting in a long position of  $\beta = 10$  shares of a stock with initial price  $S_0 = 100$ . The intra-day log return of the asset is given by  $\Delta_1 Y_{t+1} = \log(S_{t+1}/S_t)$  for  $t \ge 0$  are assumed to be iid according to a Gaussian law of mean 0 and standard deviation  $\sigma = 0.1$ .

a. Compute the VaR<sub> $\alpha$ </sub>( $L_1$ ) where  $L_1$  is the portfolio loss between today and tomorrow for  $\alpha = 99\%$ .

Hint : Use the fact that  $\Phi^{-1}(\alpha) \approx 2.3$ .

b. We keep the long position on the portfolio during 100 days. Compute  $\operatorname{VaR}_{\alpha}(L_{100})$  where  $L_{100}$  is the portfolio loss during 100 days.

## Exercise 3

Let L be a loss with law  $\mathcal{N}(\mu, \sigma^2)$ . As usual, we denote by  $\Phi$  the cdf of  $\mathcal{N}(0, 1)$  and by  $\phi$  its density. Prove that

$$\mathrm{ES}_{\alpha}(L) = \mu + \sigma \frac{\phi(\Phi^{-1}(\alpha))}{1 - \alpha}.$$

#### Exercise 4

Let  $X = (X_1, X_2) \sim \mathcal{N}(\mu, \Sigma)$  with  $\mu = (\mu_1, \mu_2)$  and

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}, \quad \rho \in (-1, 1).$$

a. Show that

$$\operatorname{Var}_{\alpha}(X_1 + X_2) \le \operatorname{Var}_{\alpha}(X_1) + \operatorname{VaR}_{\alpha}(X_2).$$

b. What's the financial interpretation of this inequality?

## Exercise 5

Prove the following relations :

a. If L follows a Laplace distribution (double exponential) with density  $f(\ell) = \frac{\lambda}{2} \exp(-\lambda |\ell|)$  then

$$\operatorname{Var}_{\alpha}(L) = \begin{cases} -\frac{1}{\lambda} \ln(2(1-\alpha)), & \text{if} \quad \alpha \ge 1/2, \\ \frac{1}{\lambda} \ln(2\alpha), & \text{otherwise} \end{cases}$$

and

$$\operatorname{ES}_{\alpha}(L) = \begin{cases} \frac{1}{\lambda}(1 - \ln(2(1 - \alpha))), & \text{if } \alpha \ge 1/2, \\ \frac{1}{\lambda}(\frac{1}{2} + \alpha(1 - \ln(2\alpha))), & \text{otherwise.} \end{cases}$$

b. If L follows a Pareto distribution with index p with density function  $f(\ell) = \frac{p}{\ell^{p+1}} \mathbf{1}_{\ell \geq 1}$ , with p > 1, then

$$\operatorname{Var}_{\alpha}(L) = (1-\alpha)^{-\frac{1}{p}}, \quad \operatorname{ES}_{\alpha}(L) = \frac{p}{p-1}(1-\alpha)^{-\frac{1}{p}}.$$

**Exercise 6** (Coherence and independence)

We could think that the risk of two independent risks aggregate together, namely, if  $L_1$ and  $L_2$  are two independent real-valued random variables and  $\rho$  is a risk measure, then

$$\rho(L_1 + L_2) = \rho(L_1) + \rho(L_2).$$

In general, this is wrong.

Exhibit two independent random variables  $L_1$  and  $L_2$  as well as a risk measure  $\rho$  such that

$$\rho(L_1 + L_2) < \rho(L_1) + \rho(L_2).$$

Exercise 7 (Conditioning a Gaussian vector)

With the notation introduced in the course, we let  $X = (X_1, X_2)$  be a Gaussian vector with mean  $\mu = (\mu_1, \mu_2)$  and covariance matrix

$$\Sigma = \begin{pmatrix} \Sigma_{1,1} & \Sigma_{1,2} \\ \Sigma_{2,1} & \Sigma_{2,2} \end{pmatrix}$$

a. Prove that

$$\mathbb{E}[X_2|X_1] = \mu_2 + \Sigma_{2,1} \Sigma_{1,1}^{-1} (X_1 - \mu_1)$$

b. Deduce from the previous question that  $X_2 - \mathbb{E}[X_2|X_1]$  is independent of  $X_2$ .

**Exercise 8** (Characteristic function of normal mixture)

- a. Recall the definition of a normal mixture with non-negative weight random variable W and parameters  $\mu \in \mathbb{R}^d$  and  $A \in \mathbb{R}^{d \times k}$ .
- b. Prove that the characteristic function of the normal mixture X is given by

$$\phi_X(u) = \mathbb{E}[\exp(iu^T X)] = \exp(iu^T \mu) F_W(\frac{1}{2}u^T \Sigma u),$$

where for  $\theta \in \mathbb{R}$ 

$$F_W(\theta) = \int_0^\infty \exp(-\theta w) \mathbb{P}_W(dw).$$

**Exercise 9** (Stochastic representation of Elliptical distributions) Prove that  $X \sim E_d(\mu, \Sigma, \psi)$  if and only if

$$X \stackrel{d}{=} \mu + RAS$$

where S is uniformly distributed on the unit sphere,  $R \ge 0$  is a radial random vector independent of S and  $A \in \mathbb{R}^{d \times k}$  such that  $\Sigma = AA^T$ .

#### Exercise 10

- a. Let  $X \sim \mathcal{N}_d(\mu, \Sigma)$ . Prove that  $X \sim E_d(\mu, \Sigma, \psi)$  for an explicit function  $\psi$ .
- b. Let  $X \sim E_d(\mu, \Sigma, \psi)$ . Prove that for any  $B \in \mathbb{R}^{d \times k}$ ,  $b \in \mathbb{R}^k$ , it holds

 $BX + b \sim E_k(B\mu + b, B\Sigma B^T, \psi).$ 

- c. Prove that if  $X = (X_1, X_2) \sim E_d(\mu, \Sigma, \psi), X_1 \in \mathbb{R}^k$  and  $X_2 \in \mathbb{R}^{d-k}$  with  $\mu = (\mu_1, \mu_2)$  and  $\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$  then  $X_1 \sim E_k(\mu_1, \Sigma_{11}, \psi), \quad X_2 \sim E_{d-k}(\mu_2, \Sigma_{22}, \psi)$
- d. Prove that if  $X \sim E_d(\mu, \Sigma, \psi)$  and  $Y \sim E_d(\tilde{\mu}, \Sigma, \tilde{\psi})$  are independent then  $X + Y \sim E_d(\mu + \tilde{\mu}, \Sigma, \psi \tilde{\psi})$ .

**Exercise 11** (sub-additivity of VaR for elliptical distributions) Let  $X \sim E_d(\mu, \Sigma, \psi)$ . Prove that for any  $u, w \in \mathbb{R}^d$  and  $\alpha \in (0, 1)$ ,  $\operatorname{VaR}_{\alpha}(u^T X + w^T X) \leq \operatorname{VaR}_{\alpha}(u^T X) + \operatorname{VaR}_{\alpha}(w^T X).$ 

Exercise 12 (Correlation bounds)

Let  $X_1, X_2$  be two log-normal distributions :  $\log(X_1) \sim \mathcal{N}(0, 1)$  and  $\log(X_2) \sim \mathcal{N}(0, \sigma^2)$  for some  $\sigma > 0$ .

- a. Compute the minimal and maximal correlation  $\rho_{\min}$ ,  $\rho_{\max}$  of the vector  $(X_1, X_2)$ .
- b. What are the limits of  $\rho_{\min}$ ,  $\rho_{\max}$  when  $\sigma \uparrow \infty$ ?

Exercise 13 (The Clayton Copula)

The Clayton copula is defined by

$$C(u_1, u_2) = \max(u_1^{-\theta} + u_2^{-\theta} - 1, 0)^{-\frac{1}{\theta}}, \quad \theta \ge -1.$$

- a. For which value of  $\theta$ , does C correspond to the minimal copula  $C_{\min}(u_1, u_2) = \max(u_1 + u_2 1, 0)$ ?
- b. Prove that when C goes to  $C_{\max}(u_1, u_2) = \min(u_1, u_2)$  when  $\theta \to +\infty$ .
- c. To which copula C converges when  $\theta \to 0\,?$